

# UNIVERSITA' DI ROMA TOR VERGATA

## Laurea Magistrale in Matematica Pura e Applicata

Corso di *EP-Elementi di Probabilità*

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**Exercise 1** Let  $B$  a Brownian motion and let, for  $t \geq 0$ ,

$$Z_t = \frac{1}{\sqrt{1+t}} B_{1+t}.$$

- a) Is it a Markov process? Which is its transition function? Which is its initial distribution?
- b) Is it time homogeneous? Prove that it is not a stationary process.
- c) Let  $(\Omega, \mathcal{F}, (\mathcal{F}_s)_s, (Z_t)_t, (\mathbf{P}^{x,s})_{x,s})$  a realization of the Markov process associated to the transition function computed in a). Prove that, whatever the starting point  $x$  and the initial time  $s$ ,

$$Z_t \xrightarrow[t \rightarrow +\infty]{} N(0, 1) \quad \text{in law}.$$

**Exercise 2** Let  $B = (\Omega, \mathcal{F}, (\mathcal{F}_t)_t, (B_t)_t, \mathbf{P})$  be a  $m$ -dimensional Brownian motion and let  $X_t = B_{ct}$ , where  $c > 0$ .

- a) Prove that  $X$  is a time-homogeneous Markov process and determine its transition function  $p$ .
- b1) Let  $X = (\Omega, \mathcal{F}, (\mathcal{F}_t)_t, (X_t)_t, (\mathbf{P}^{x,s})_x)$  the Markov process associated to function  $p$ . Prove that the distribution of  $X_t$  under  $\mathbf{P}^{x,s}$  is the same as the distribution of  $x + B_{c(t-s)}$ .
- b2) Which is the generator of  $X$ ?

**Exercise 3** Let  $X = (\Omega, \mathcal{F}, (\mathcal{F}_t)_t, (X_t)_t, \mathbf{P})$  be a time homogeneous Markov process taking values in some measurable space  $(E, \mathcal{E})$  and associated to a transition function  $p$ . Let  $A \in \mathcal{E}$ .

- a) For  $s \leq t$  compute

$$\mathbf{E}[1_A(X_t) | \mathcal{F}_s]$$

and express it as a function of  $s$ ,  $X_s$  and  $p$ .

- b) Let, for  $s \leq t$ ,

$$M_s = p(t - s, X_s, A).$$

- b1) Prove that  $(M_s)_{s \leq t}$  is a  $(\mathcal{F}_t)_{s \leq t}$ -martingale. Is it uniformly integrable?
- c) Assume in addition that  $X$  is a Brownian motion.
- c1) Write explicitly the expression of  $p(t - s, x, A)$ . Remark that for  $s < t$  this is a continuous function of  $(s, x)$  (just give an argument, I do not ask for a complete proof). Deduce that  $M$  is a continuous martingale. Prove that it converges a.s. as  $s \nearrow t$ .
- c2) Let  $\mathcal{F}_{t-} = \bigvee_{s < t} \mathcal{F}_s$  the  $\sigma$ -algebra generated by  $\bigcup_{s < t} \mathcal{F}_s$ . Let  $M_\infty = \lim_{s \uparrow t} M_s$ . Prove that  $M_\infty = \mathbf{E}[1_A(X_t) | \mathcal{F}_{t-}]$  (Hint: recall the proof of Proposition 4.21).
- c3) Assume that  $(\mathcal{F}_t)_t$  is augmented (i.e. that for every  $t$   $\mathcal{F}_t$  contains the negligible events of  $\mathcal{F}$ ). Prove that  $\lim_{s \uparrow t} M_s = 1_A(X_t)$  a.s.

## Solutions

**Exercise 1.** a) The covariance function of  $Z$  is, for  $s \leq t$ ,

$$K_{t,s} = \text{Cov}(X_s, X_t) = \frac{1}{\sqrt{1+t}\sqrt{1+s}} \mathbb{E}[B_{1+t}B_{1+s}] = \frac{1+s}{\sqrt{1+t}\sqrt{1+s}} = \frac{\sqrt{1+s}}{\sqrt{1+t}}$$

and we have, for  $u \leq s \leq t$

$$K_{u,t} = \frac{\sqrt{1+u}}{\sqrt{1+t}},$$

$$K_{u,s} K_{s,s}^{-1} K_{s,t} = \frac{\sqrt{1+s}}{\sqrt{1+t}} \frac{\sqrt{1+u}}{\sqrt{1+s}} = \frac{\sqrt{1+u}}{\sqrt{1+t}}$$

so that this is a Markov process. Its transition function  $p(s, t, x, dy)$ ,  $s \leq t$ , is the conditional distribution of  $X_t$  given  $X_s = x$ , which is of course Gaussian with mean

$$\mathbb{E}(X_t) + \frac{K_{t,s}}{K_{s,s}} (x - \mathbb{E}(X_s)) = \frac{\sqrt{1+s}}{\sqrt{1+t}} x$$

and variance

$$\text{Var}(X_t) - \frac{K_{t,s}^2}{K_{s,s}} = 1 - \frac{1+s}{1+t} = \frac{t-s}{1+t}$$

i.e.  $p(s, t, x, dy) \sim N(\frac{\sqrt{1+s}}{\sqrt{1+t}} x, \frac{t-s}{1+t})$ . The initial distribution is the distribution of  $B_1 \sim N(0, 1)$ .

b) As the law  $p(s, t, x, dy)$  does depend on the couple  $s, t$  and not just on the difference  $t - s$ ,  $Z$  is not time homogeneous and *a fortiori* is not stationary (the joint distributions of  $(X_s, X_t)$  and  $(X_{s+h}, X_{t+h})$  are different as they have different covariance matrices).

c) The law of  $Z_t$  with starting point  $x$  and initial time  $s$  is  $p(s, t, x, dy) \sim N(\frac{\sqrt{1+s}}{\sqrt{1+t}} x, \frac{t-s}{1+t})$ . As for  $t \rightarrow +\infty$  the mean converges to 0 and the variance to 1,  $Z_t$  converges in law to a  $N(0, 1)$  distribution for every initial conditions  $x, s$ .

**Exercise ‘monique1’.** a) We have

$$\mathbb{E}[X_t | \mathcal{F}_s] = \mathbb{E}[ab + (b-a)B_t - B_t^2 + t | \mathcal{F}_s] = ab - (b-a)B_s - \mathbb{E}[(B_t - B_s) + B_s]^2 | \mathcal{F}_s]$$

Now

$$\begin{aligned} \mathbb{E}[(B_t - B_s) + B_s]^2 | \mathcal{F}_s] &= \mathbb{E}[(B_t - B_s)^2 + B_s^2 - 2B_s(B_t - B_s) | \mathcal{F}_s] = \\ &= t - s + B_s^2 - 2B_s \underbrace{\mathbb{E}[B_t - B_s | \mathcal{F}_s]}_{=0} = t - s + B_s^2 \end{aligned}$$

and substituting we find  $\mathbb{E}[X_t | \mathcal{F}_s] = X_s$ .

b) First of all remark that  $\tau$  is finite, thanks to the Iterated Logarithm Law. Then by the stopping theorem we have, for every  $n > 0$ ,

$$ab = E(X_0) = E(X_{\tau \wedge n}) = (b - B_{\tau \wedge n})(a + B_{\tau \wedge n}) + E(\tau \wedge n)$$

Now the quantity  $(b - B_{\tau \wedge n})(a + B_{\tau \wedge n})$  is bounded in  $n$  as  $-a \leq B_{\tau \wedge n} \leq b$ . As  $B_{\tau \wedge n} \rightarrow B_\tau$  as  $n \rightarrow \infty$  and  $B_\tau$  can only take the values  $-a$  or  $b$ , we have  $(b - B_{\tau \wedge n})(a + B_{\tau \wedge n}) \rightarrow 0$  as  $n \rightarrow \infty$  and also  $E[(b - B_{\tau \wedge n})(a + B_{\tau \wedge n})] \rightarrow 0$  by Lebesgue's theorem. Finally  $E(\tau \wedge n) \rightarrow E(\tau)$  by Beppo Levy's theorem so that we have

$$E(\tau) = ab .$$

**Exercise 2.** a)  $X$  being clearly a Gaussian process we just need to check the relation  $K_{u,t} = K_{u,s} K_{s,s}^{-1} K_{s,t}$  for  $u \leq s \leq t$ , which is immediate as, for  $s \leq t$ ,

$$K_{s,t} = \text{Cov}(B_{cs}, B_{ct}) = cs .$$

The transition probability  $p(t - s, x, dy)$  is given by the conditional distribution of  $X_t$  given  $X_s = x$ , which is Gaussian with mean

$$E(X_t) + \frac{K_{s,t}}{K_{s,s}}(x - E(X_s)) = x$$

and variance

$$K_{t,t} - \frac{K_{s,t}^2}{K_{s,s}} = c(t - s)$$

i.e.  $p(t, x, dy) \sim N(x, ct)$ .

b1) This is immediate, as the distribution of  $X_t$  starting at  $x$  is  $p(t, x, dy)$ , which was just computed and that is a  $N(x, ct)$  distribution, that coincides with the distribution of  $x + B_{ct}$ .

b2) We must compute

$$Lf(x) = \lim_{h \rightarrow 0} \frac{1}{h} (E^x[f(X_h)] - f(x)) = \lim_{h \rightarrow 0} \frac{1}{h} (E[f(x + B_{ch})] - f(x))$$

We know that, if  $f$  is bounded and twice differentiable,

$$\lim_{h \rightarrow 0} \frac{1}{h} (E[f(x + B_h)] - f(x)) = \frac{1}{2} \Delta f(x)$$

(the generator of the Brownian motion). Therefore, for every  $f$  bounded and twice differentiable,

$$Lf(x) = \frac{c}{2} \Delta f(x)$$

**Exercise 3.** a) By the Markov property

$$(1) \quad E[1_A(X_t) | \mathcal{F}_s] = p(t-s, X_s, A) .$$

b1) From a) we know that  $M_s = E[1_A(X_t) | \mathcal{F}_s]$ , so that the martingale property is immediate. Also uniform integrability is immediate from (1) and Proposition 4.17.

c1) We have

$$p(t-s, x, A) = \frac{1}{\sqrt{2\pi(t-s)}} \int_A \exp\left(-\frac{(y-x)^2}{2(t-s)}\right) dy$$

and the continuity in  $(s, x)$  follows from the usual results about integrals depending on a parameter. As a consequence  $s \mapsto M_s$  turns out to be a composition of continuous functions, thus continuous itself. The existence of the a.s. and  $L^1$  limit is a consequence of Theorem 4.24, as seen in class.

c2) We must prove that we have

$$(2) \quad E[M_\infty 1_G] = E[1_A(X_t) 1_G]$$

for every set  $G \in \mathcal{F}_{t-}$ , or at least for  $G$  in a family  $\mathcal{C} \subset \mathcal{F}_{t-}$  that is stable with respect to finite intersections and generating  $\mathcal{F}_{t-}$ . A convenient family is  $\mathcal{C} = \bigcup_{s < t} \mathcal{F}_s$ . Now if  $G \in \mathcal{C}$ , then  $G \in \mathcal{F}_s$  for some  $s$  and

$$E[M_\infty 1_G] = \lim_{u \rightarrow t} E[E[1_A(X_t) | \mathcal{F}_u] 1_G] = \lim_{u \rightarrow t} E[E[1_A(X_t) 1_G | \mathcal{F}_u]] = E[1_A(X_t) 1_G]$$

where we have used the fact that  $E[1_A(X_t) | \mathcal{F}_u] 1_G = E[1_A(X_t) 1_G | \mathcal{F}_u]$ , as soon as  $u \geq s$ . Therefore (2) is proved.

c3) This would be immediate from c2) if we knew that the r.v.  $1_A(X_t)$  is  $\mathcal{F}_{t-}$ -measurable (by now we only know that it is  $\mathcal{F}_t$ -measurable).

But the filtration  $(\mathcal{F}_s)_s$  certainly contains the natural filtration  $(\mathcal{G}_s)_s$  of the Brownian motion. Hence also the augmented natural filtration  $(\overline{\mathcal{G}}_s)_s$ . Therefore, finally, just remark that  $\mathcal{F}_{t-}$  contains  $\overline{\mathcal{G}}_{t-}$  which coincides with  $\overline{\mathcal{G}}_t$  by the fundamental property of continuity of the augmented natural filtration (Proposition 3.3). As  $1_A(X_t)$  is  $\overline{\mathcal{G}}_t$ -measurable, then it is also  $\mathcal{F}_{t-}$ -measurable which allows to conclude.